

Normal correlation analysis

Part 5 of regression

$$f(x,y) = \frac{1}{2\pi \sigma_x \sigma_y \sqrt{1-\rho^2}} e^{-\frac{1}{2(1-\rho^2)} \left[\left(\frac{x-\mu_x}{\sigma_x}\right)^2 + \left(\frac{y-\mu_y}{\sigma_y}\right)^2 - 2\rho \left(\frac{x-\mu_x}{\sigma_x}\right) \left(\frac{y-\mu_y}{\sigma_y}\right) \right]}$$

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$$f(x,y; \sigma_x^2, \sigma_y^2, \rho, \mu_x, \mu_y)$$

$$(x_1, y_1) \quad (x_2, y_2) \quad \dots \quad (x_n, y_n)$$

use MLE to estimate values of parameters
 $\mu_x, \mu_y, \sigma_x^2, \sigma_y^2, \rho$

$$L(\sigma_x^2, \sigma_y^2, \rho, \mu_x, \mu_y; \bar{x}, \bar{y}) = \prod_{k=1}^n f(x_k, y_k; \sigma_x^2, \sigma_y^2, \rho, \mu_x, \mu_y)$$

$$\text{maximize } n \ln L \quad \frac{\partial}{\partial \rho} \ln L = 0$$

$$\frac{\partial}{\partial \mu_x} \ln L = 0 \quad \dots$$

$$\mu_x = \bar{x} \quad \sigma_x^2 = \frac{1}{n} \sum (x_i - \bar{x})^2 = s_{xx}$$

$$\mu_y = \bar{y} \quad \sigma_y^2 = \frac{1}{n} \sum (y_i - \bar{y})^2 = s_{yy}$$

$$\sigma_{xy} = \frac{1}{n} \sum (x_i - \bar{x})(y_i - \bar{y}) = s_{xy}$$

$$\rho = \frac{S_{xy}}{\sqrt{S_{xx}S_{yy}}}$$

get estimates $\hat{\mu}_x = \bar{x}$ $\hat{\mu}_y = \bar{y}$ $\hat{\sigma}_x^2 = S_{xx}$
 $\hat{\sigma}_y^2 = S_{yy}$

$$\hat{\rho} = \frac{S_{xy}}{\sqrt{S_{xx}S_{yy}}}$$

Ex: want interval estimate for σ_x^2 $\left(\frac{\hat{\sigma}_x^2}{\sigma_x^2} \chi_{n-1}^2 \right)$
 using χ^2 , t variables, can get interval estimates
 can do hypothesis testing for values of
 $\sigma_x^2, \sigma_y^2, \mu_x, \mu_y$

$$\hat{\rho} = \frac{S_{xy}}{\sqrt{S_{xx}S_{yy}}} \quad \text{sample value for this is referred to by "r"}$$

$$\sigma_{y|x}^2 = \sigma_y^2 (1 - \rho^2) \quad -1 \leq \rho \leq 1 \quad \rho=0 \quad (x_i, y_i \text{ indep})$$

$$\left. \quad \quad \quad \right\} \quad \rho = \pm 1$$

$$\rho^2 = \frac{\sigma_y^2 - \sigma_{y|x}^2}{\sigma_y^2} = 1 - \left(\frac{\sigma_{y|x}^2}{\sigma_y^2} \right)$$

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fraction of variance
 "due to x"
Special case: if $\rho = 0$

the fraction of variance
 remains after x is known

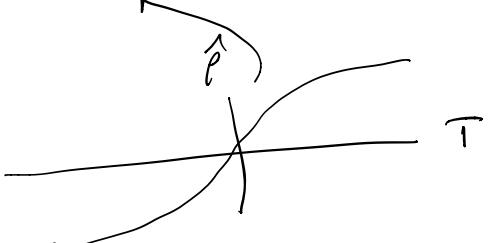
$$\hat{p} \quad T = \hat{p} \sqrt{\frac{n-2}{1-\hat{p}^2}}$$

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turns out this
 is a t-distributed
 variable w/
 n-2 degrees
 of freedom

$$\hat{p} = \frac{T}{\sqrt{n-2+T^2}}$$

mean. memory add
 1-1 function



$$P(\hat{p} > \lambda \mid \rho = 0)$$

$$\stackrel{\uparrow}{=} P\left(\hat{p} \sqrt{\frac{n-2}{1-\hat{p}^2}} > \lambda \sqrt{\frac{n-2}{1-\lambda^2}}\right)$$

$$\frac{1}{T} P(T > \lambda \sqrt{\frac{n-2}{1-\lambda}})$$

What if $\rho \neq 0$? $\hat{\rho}$

$$P(\hat{\rho} \geq 0.7 | \rho = 0.5) ?$$

Can write an exact expression for density func. for $\hat{\rho}$

$$f(r) = \frac{(n-2) \Gamma(n-1) (1-\rho^2)^{\frac{n-1}{2}} (1-r^2)^{\frac{n-4}{2}}}{\sqrt{2\pi} \Gamma(n-\frac{1}{2}) (1-\rho r)^{n-\frac{3}{2}}} \cdot {}_2F_1\left(\frac{1}{2}, \frac{1}{2}; \frac{2n-1}{2}; \frac{\rho r + 1}{2}\right)$$

$${}_2F_1(a, b; c)x = \sum_{k=0}^{\infty} \frac{a^k b^k}{c^k} \frac{x^k}{k!}$$

$$m^k = m(m+1)(m+2)\dots(m+k-1)$$

if $k \geq 1$

$$m^0 = 1$$