

Partial model

$$f(x) = \lambda e^{-\lambda x}$$

λ unknown \rightarrow

$$= \frac{1}{\sqrt{2\pi\sigma^2}} e^{-\frac{1}{2}(x-\mu)^2/\sigma^2}$$

μ, σ^2 unknown \rightarrow

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$$f(x) = f_{\theta}(x) = f(x; \theta) = f(x, \theta)$$

make observations

estimate θ
 θ 's

more specific model
some confidence intervals
usually make θ 's.

Specific examples

$$\begin{array}{l} \mu \rightsquigarrow \bar{X} \\ \sigma^2 \rightsquigarrow S^2 \end{array}$$

Want: general procedures which take some general form of our distribution of unknown parameters \rightsquigarrow estimators

This week:

- Method of moments
 - Method of Maximum Likelihood.
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Method of moments

Def k^{th} sample moment (about 0)

$$M_k^i = \frac{1}{n} \sum_{i=1}^n X_i^k \quad \text{random var}$$

$$m_k^i = \frac{1}{n} \sum_{i=1}^n x_i^k \quad \text{observers.}$$

X_i iid rand. var.
 $f_X(x)$

Def $\mu_k^i = E[X^k]$ $X =$ rand var w/ dist. of pop

M_k^i are unbiased estimators of μ_k^i

Given exp. population: $f_X(x) = \lambda e^{-\lambda x}$

Method of moments: solve for unknown parameters in terms of moments, solve gives estimators for parameters

$$\mu = \frac{1}{\lambda} \quad \text{know } \mu_1' = \mu$$

$$\text{have } M_1' = \bar{X} \quad \text{estimate for } \mu$$

$$\Rightarrow \frac{1}{M_1'} = \frac{1}{\bar{X}} \quad \text{estimate for } \lambda$$

↗
biased.

$$f_{\lambda}(x) = \lambda e^{-\lambda x} \quad f_{\theta}(x) = \frac{1}{\theta} e^{-x/\theta}$$

$$\theta = \frac{1}{\lambda} \quad \mu = \theta \quad \bar{X} \text{ estimate for } \theta$$

↗
unbiased ↗ μ

Example:

Gamma variable:

$$f_{\alpha, \beta}(x) = \frac{1}{\beta^{\alpha} \Gamma(\alpha)} x^{\alpha-1} e^{-x/\beta} \quad x > 0$$

$$\mu_1' = \alpha\beta \quad \mu_2' = \alpha(\alpha+1)\beta^2$$

" $E[x^2]$

estimates for α, β

$$m_1' = \frac{1}{n} \sum X_i \quad m_2' = \frac{1}{n} \sum X_i^2$$

$\left. \vphantom{\sum X_i} \right\} \mu_1'$
 $\left. \vphantom{\sum X_i^2} \right\} \mu_2'$

$$\mu_1' = \alpha \beta \quad \alpha = \frac{\mu_1'}{\beta} \quad \mu_2' = \alpha(\alpha+1)\beta^2$$

$$\mu_2' = \frac{\mu_1'}{\beta} \left(\frac{\mu_1'}{\beta} + 1 \right) \beta^2$$

$$\mu_2' = \left(\frac{(\mu_1')^2}{\beta^2} + \frac{\mu_1'}{\beta} \right) \beta^2$$

$$\mu_2' = (\mu_1')^2 + \mu_1' \beta$$

$$\beta = \frac{\mu_2' - (\mu_1')^2}{\mu_1'}$$

similarly $\alpha = \frac{(\mu_1')^2}{(\mu_2') - (\mu_1')^2}$

$$\hat{\beta} = \frac{M_2' - (M_1')^2}{M_1'} \quad \hat{\alpha} = \frac{(M_1')^2}{M_2' - (M_1')^2}$$

For side topic:

why was

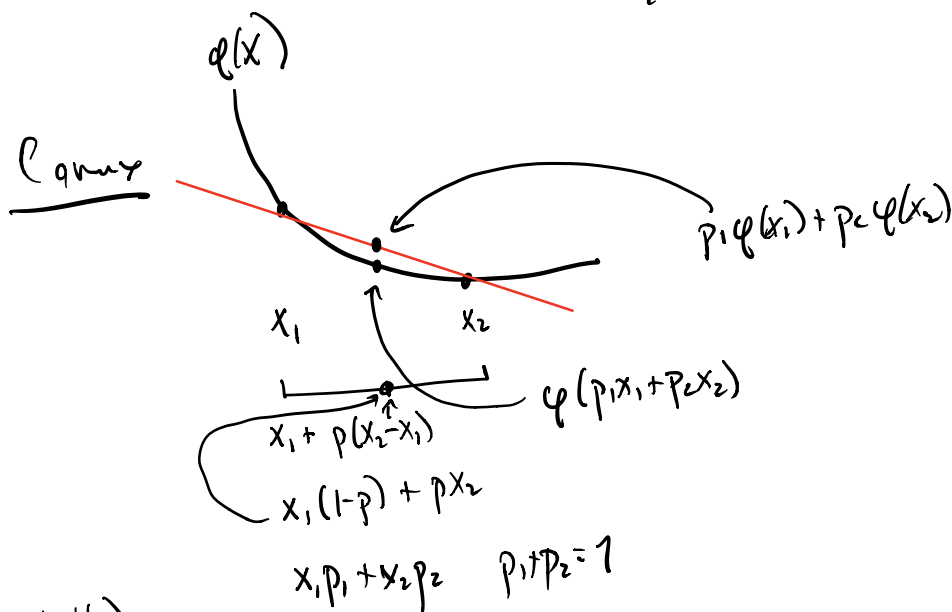
$\frac{1}{\bar{x}}$ biased estimator for $\lambda = \frac{1}{\mu}$

Jensen's inequality

X is a random variable, φ convex function then

$$\varphi(E(X)) \leq E[\varphi(X)]$$

with equality if φ is linear or in case φ is strictly convex, only if $\text{Var}(X) = 0$.



Def φ ^(strictly) convex if for $x_1 < x_2$

$$\varphi(p_1x_1 + p_2x_2) \leq p_1\varphi(x_1) + p_2\varphi(x_2)$$

($\hat{<}$)